

An affine intensity model for large credit portfolios

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- ▶ We model the intensity of default of N obligors.
- ▶ We assume that the obligors belong to S different groups (industrial sectors) and to a unique general environment (market).
- ▶ The default of any obligor can be due to one of the following causes (credit events):
 1. idiosyncratic
 2. sectoral
 3. general

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The model

- ▶ The time of default τ_j of obligor j has intensity

$$\lambda_j = X_j + u_j Y_{g(j)} + v_j Z, \quad j = 1, \dots, N,$$

where $\{X_1, \dots, X_N, Y_1, \dots, Y_S, Z\}$ are independent affine processes. X_1, \dots, X_N are the idiosyncratic factors, Y_1, \dots, Y_S are the sectoral factors Z is the common factor.

- ▶ The factors are affine jump-diffusions,

$$dX(t) = k(\theta - X(t))dt + \sigma\sqrt{X(t)}dW(t) + dJ(t),$$

where W is Brownian motion, J a pure-jump process with jump sizes exponential with mean μ , and jump times independent Poisson process with arrival rate ℓ .

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- ▶ **Note:** λ_j in general is not affine.

Related literature

- ▶ We generalize a model by Duffie and Gârleanu (2001) for the valuation of CDO's. DG restrict λ_j 's to be affine, consider only one common factor, imposing to all the obligors the same sensibility to it.
- ▶ DG illustrate the effect of correlation and prioritization on CDO's valuation
- ▶ A common critique: intensity models produce low correlations of defaults (Hull and White (2001))
- ▶ Yu (2002) provides some evidence that this is not always true
- ▶ Other affine credit models include Duffee (1999) and Driessen (2002). They focus on the identification of default risk premium on corporate bonds.

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Survival probabilities

► The survival probability

$$\begin{aligned}
 \mathbb{P}(\tau_j > t) &= \mathbb{E}_0 e^{-\int_0^t \lambda_j(u) du} \\
 &= \mathbb{E}_0 e^{-\int_0^t X_j(u) du} e^{-u_j \int_0^t Y_{g(j)}(u) du} e^{-v_j \int_0^t Z(u) du} \\
 &= \mathbb{E}_0 e^{-\int_0^t X_j(u) du} \mathbb{E}_0 e^{-u_j \int_0^t Y_{g(j)}(u) du} \mathbb{E}_0 e^{-v_j \int_0^t Z(u) du}
 \end{aligned}$$

... because the processes are affine:

$$f(X, t, m) := \mathbb{E}_0 e^{-\int_0^t m X(u) du} = \exp(C(X, t, m)),$$

where

$$C(X, t, m) = \alpha_\psi^m(t) + \beta_\psi^m(t) X(0).$$

The functions $\alpha_\psi^m(t), \beta_\psi^m(t)$ can be efficiently computed.

Factorization of survival probabilities

- ▶ The survival probabilities of the i -th obligor can be factorized as

$$s_i^t := f(X_i, t, 1)f(Y_{g(i)}, t, u_i)f(Z, t, v_i),$$

The formula shows the dependence of the default probability of a single obligor on the occurrences of idiosyncratic, sectoral, general "credit events".

- ▶ Let $e(M, t)$ be the r.v. indicating the occurrence of a general credit event before time t , then

$$\mathbb{P}(e(M, t)) = 1 - f(Z, t, 1)$$

and

$$\mathbb{P}(\tau^i < t, e(M, t)) = 1 - f(Z, t, v_i)$$

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Data needed for calibration

- ▶ For each obligor i , the survival probabilities

$$s_i^t := \mathbb{P}(\tau_i > t), \quad i = 1, \dots, N, \quad t \in \mathcal{T}.$$

- ▶ The probabilities of general and sectoral credit events

$$p^M(t) := \mathbb{P}(e(M, t))$$

$$p^i(t) := \mathbb{P}(e(G(i), t)), \quad i = 1, \dots, S$$

- ▶ For each obligor i , the probability of default conditioned on a general or sectoral credit event

$$p_{i|G}(t) := \mathbb{P}(\tau^i < t | e(G(i), t))$$

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Calibration procedure

- ▶ Estimate the parameters sets $\hat{\psi}_c$ and $\{\hat{\psi}_{g_i}, i = 1, \dots, S\}$ that better fit the input data $\{p^M(t), t \in \mathcal{T}\}$ and $\{p^i(t), i = 1, \dots, S, t \in \mathcal{T}\}$
- ▶ From $p_{i|G}(t)$ and $p^i(t)$, extract $\mathbb{P}(\tau^i < t, e(G(i), t))$ and compute u_i from

$$u_i := \operatorname{argmin}_u \|\mathbb{P}(\tau^i < t, e(G(i), t)) - (1 - f(Y_{g_i}, t, u))\|$$

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- ▶ Fit the obligors' survival probabilities s_i^t by calibrating the parameters of the idiosyncratic factor X_i




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An example of calibration

- ▶ An exercise with default probabilities taken from Moody's transition matrix 1980-1999.
- ▶ We consider a portfolio of 90 obligors, belonging to 3 groups.
- ▶ Within each group, obligors 1-10 are Ba, 11-20 are B, 21-30 are Caa 
- ▶ General credit events are rare (Aa) 
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- ▶ The dependence is only considered at time $t = 1$
- ▶ We assume three levels of dependence
 $L = 0.25$; $M = 0.5$; $H = 0.75$, that is

$$p_{i|G}(t) = L, M, H$$

- ▶ Within each group, obligors 1-10 have dependence (with their group and with the general factor) L , 11-20 M , 21-30 H .

▶ Plot

- ▶ We can compute the correlation of defaults before time t

$$\rho_{i,j} = \text{corr}(\mathbf{1}_{\tau^i < t}, \mathbf{1}_{\tau^j < t})$$

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Simulations

The standard way to simulate default scenarios $\{\tau^{(n)}, I^{(n)}\}_{n=1}^{N_d}$ is

- ▶ From $\lambda_j(t_0), j = 1, \dots, N$ compute the total intensity $\Lambda(t_0) = \sum_{j=1}^N \lambda_j(t_0)$
- ▶ Simulate the occurrence of one default in the interval Δt with probability $\Lambda(t_0)\Delta t$
- ▶ In case of default extract the identity of defaulter with probabilities $p_j = \lambda_j(t_0)/\Lambda(t_0)$
- ▶ Compute $\lambda_j(t_0 + \Delta t), j = 1, \dots, N$ by discretization

Simulations

- ▶ The standard way to simulate is based on the discretization of the continuous process. It converges as $\Delta t \rightarrow 0$
- ▶ The smaller Δt the greater the computational effort
- ▶ Another problem is that the discretization of affine-square root processes often produces negative values.

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Exact simulations

Exact simulation for default scenarios $\{\tau^{(n)}, I^{(n)}\}_{n=1}^{N_d}$

- ▶ I set $\tau^{(0)} := 0$, $C_0 := \{1, \dots, N\}$ and $n := 1$;
- ▶ II simulate $\tau^{(n)}$, IF $\tau^{(n)} > T$ or $n = N$ STOP;
- ▶ III extract $I^{(n)}$;
- ▶ IV re-start the intensities λ_j of obligors $j \in C_n := C_{n-1} \setminus \{I^{(n)}\}$;
- ▶ V set $n = n + 1$, GO TO (II).

A comparison of the two methods

- ▶ We simulated default scenarios from the previous calibration exercise
- ▶ 10 simulations for defaults before time $T = 7$
- ▶ Consider the discretization method. Let N be the number of intervals in a year.
- ▶ Compute the mean frequency of negative values as a function of N [4/10]
- ▶ Compute the CPU times for the 10 simulations and compare to the exact method [4/10]

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Conclusions

- ▶ A model for large portfolios subject to credit risk (e.g. CDO)
- ▶ The affine setting leads to closed formulas for many important quantities
- ▶ Flexible enough for calibration
- ▶ Can produce high correlations between defaulters
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II: Simulate the n^{th} default-time $\tau^{(n)}$

- ▶ Invert the total survival probability

$$\begin{aligned} SP_n(t) &:= \mathbb{P}(s_n > t | \mathcal{F}_{\tau^{(n-1)}}) \\ &= \mathbb{E}_{\tau^{(n-1)}} \left[e^{-\int_{\tau^{(n-1)}}^{\tau^{(n-1)}+t} \Lambda^{n-1}(u) du} \right] \end{aligned}$$

where

$$\Lambda^{n-1}(t) = \sum_{j \in C_{n-1}} \lambda_j(t) = \sum_{j \in C_{n-1}} X_j(t) + u_j Y_{g(j)}(t) + v_j Z(t)$$

- ▶ Easy to be computed and inverted (numerically).

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III: extract the n^{th} defaulter

- ▶ The probability, conditioned to $\tau^{(n)}$, that $I^{(n)} = j$ is

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IV: re-start the intensities

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$$\mathbb{E}[X_j(\tau^{(1)}) | \tau^{(1)} = t, I^{(1)} = k] = \left. \frac{\partial}{\partial r} \right|_{r=0} H^j(r; t, k), \quad j = 1, \dots, N$$



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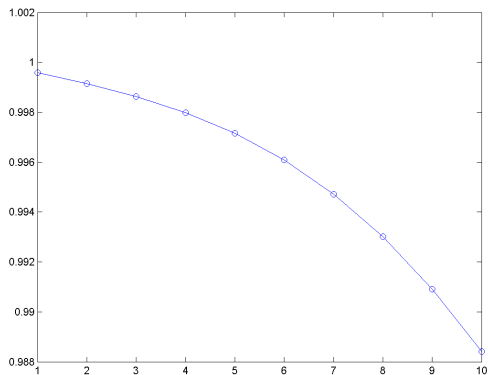
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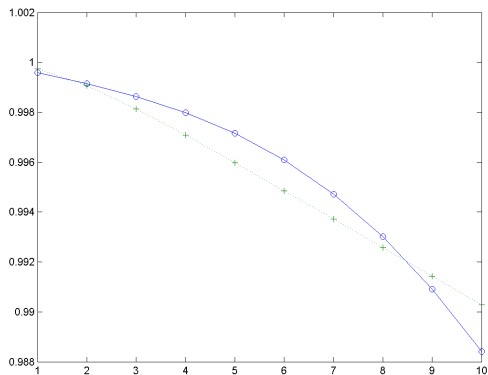
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Input data: General factor (Aa)

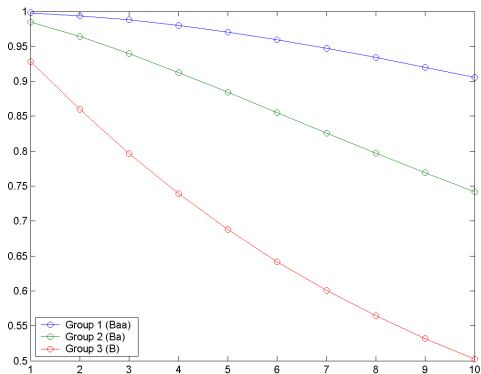
[Text](#)[Fitting](#)

Calibrated data: General factor (A_a)

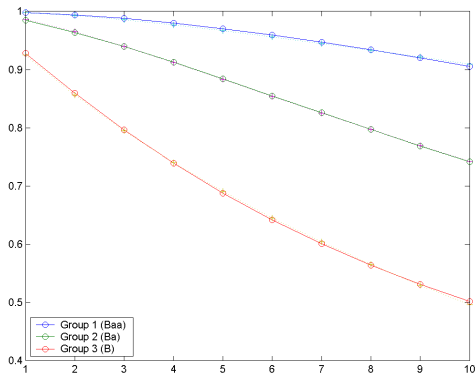


▶ Text

Input data: Groups

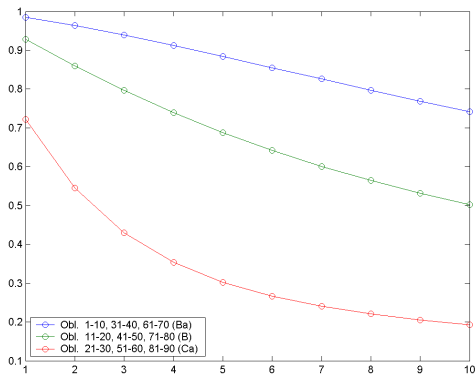
[Text](#)[Fitting](#)

Calibrated data: Groups

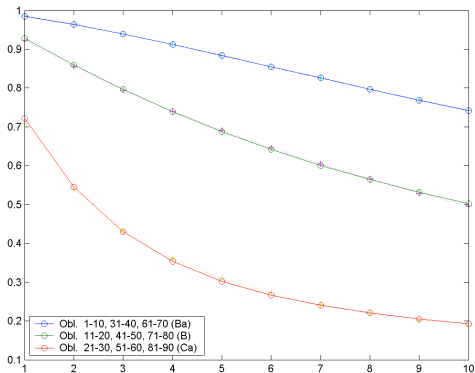


Text

Input data: obligors

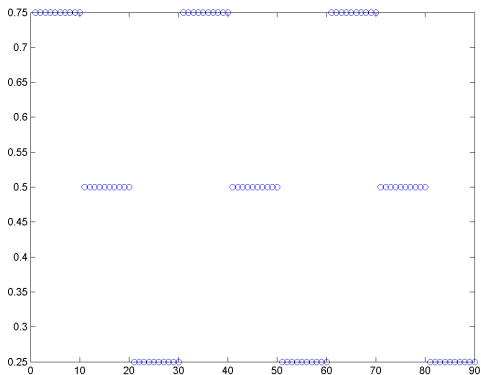
[Text](#)[Fitting](#)

Calibrated data: obligors



▶ Text

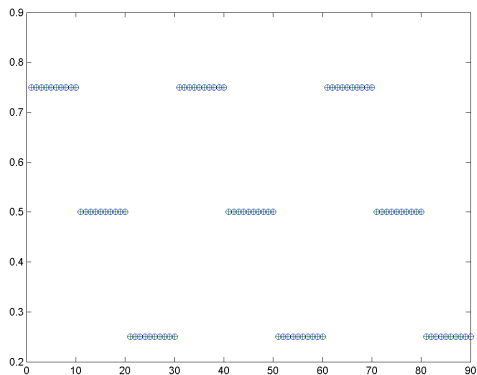
Input data: conditional probabilities (general and sectoral)



▶ Text

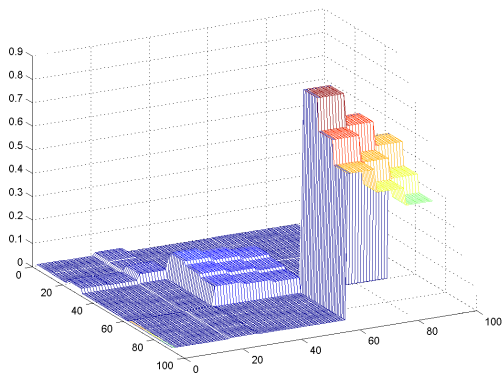
▶ Fitting

Cal. data: conditional probabilities (general and sectoral)



Text

Correlation of defaults in the first year



Text

